

BIDDERS' CHOICE AUCTIONS: RAISING REVENUES THROUGH THE RIGHT TO CHOOSE

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Abstract

Sales of multiple real-estate properties are often conducted via a sequence of ascending auctions, giving the winner at each stage the right to choose one of the available lots. We show that when bidders are risk averse, such "bidders' choice" auctions raise more revenues than standard simultaneous or sequential ascending auctions. We also report the results of laboratory experiments to investigate the effectiveness of bidders' choice auctions vis-a-vis the simultaneous ascending auction. The revenue-superiority of the bidders' choice auction is corroborated by the experimental data. Finally, we compare observed bidding behavior in the experiments with theoretically predicted bids to estimate a common risk aversion parameter from the data. (JEL: D44, C72)

1. Introduction

In its glossary of auction terms, the National Association of Realtors defines a bidders' choice auction as:¹

A method of sale whereby the successful high bidder wins the right to choose a property (or properties) from a grouping of similar or like-kind properties. After the high bidder's selection, the property is deleted from the group, and the second round of bidding commences, with the high bidder in round two choosing a property, which is then deleted from the group and so on, until all properties are sold.

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1. See http://www.aaauctionservice.com/glossery_files/glossery.htm).

According to the web site of another real-estate auctioneer, "this type of bidding is very popular when selling building lots, time-shares, and condominiums,"² see also Ashenfelter and Genesove (1992). In this paper, we investigate why sellers may wish to employ bidders' choice or "right-to-choose" auctions. The format we consider is the ascending right to choose (A-RTC) auction, which is commonly employed in the sale of real estate. The A-RTC auction has several stages of ascending bid auctions, with the high bidder at each stage winning the right to choose one item from the set of remaining items.

One interesting feature of the A-RTC auction is that it can create competition between bidders who are interested in different items for sale. Consider, for example, the case with two bidders and two items. Suppose each bidder is interested in one item only, and both are equally likely to prefer either item. When a standard simultaneous ascending (SA) auction is conducted the seller's revenue is zero when bidders prefer different items, which occurs with probability one-half. In contrast, in the first-stage of the A-RTC auction there is always competition since bidders, who do not know their rival's preference, run the risk that their preferred item is taken by the first-stage winner. The seller's revenue from an A-RTC auction is therefore strictly positive even when bidders turn out to prefer different items *ex post*.

The ability of the A-RTC auction to create competition between bidders interested in different items does not imply, however, that its expected revenue exceeds that of the SA auction. In fact, in those cases where bidders prefer the same item, they bid more aggressively in the SA auction since there is no chance of winning their preferred item in a later stage. We show below that, on average, both formats yield the same revenue when bidders are risk neutral. Hence, the widespread use of bidders' choice auctions to sell different real-estate properties, for instance, remains unexplained when bidders and the seller are risk neutral.

However, the seller's revenue from an A-RTC auction may exceed that of a SA auction when bidders are risk averse. Burguet (1999) makes this point nicely with the following example. Suppose that in the two-bidder example above, bidders' values are 1 for the preferred item and 0 for the other item. If the items are sold in the SA auction, the seller's revenue is 1 if both bidders prefer the same item and is zero otherwise. The seller's expected revenue is therefore one-half independent of bidders' risk attitudes. Now consider the first stage of an A-RTC auction. The first bidder to drop out will win the second auction at a price of zero, and therefore obtains an expected payoff of $(1/2)u(1) + (1/2)u(0)$. Hence, a bidder will remain in the first stage auction until the bid reaches the price R which makes her indifferent between winning the first stage at price R or dropping out, that is, until

2. See http://www.paxauctions.com/buyers/auction_types.html.

$$u(1 - R) = (1/2)u(1) + (1/2)u(0).$$

Since the second-stage revenue is zero, the total revenue of the A-RTC auction is R . It is straightforward to verify that R is greater than $1/2$ when $u(\cdot)$ is concave, that is, when bidders are risk averse.

Below we report the results of an experiment in which four bidders compete for two items in a SA or A-RTC auction. Bidders were equally likely to prefer either item, as in the example above. Our experimental results show that, consistent with the theoretical predictions for risk-averse bidders, the A-RTC auction generates significantly more revenue than a SA auction. From the data we estimate a common coefficient of constant relative risk aversion and find that it is $r = 0.61$. With this coefficient of risk aversion, the theoretical model of the A-RTC auction explains the data well. In particular, the empirical distribution of seller revenues obtained in our experiment closely lines up with the distribution of seller revenue predicted by the theory.

1.1. Related Literature

Burguet (1999, 2002) studies the A-RTC auction when each bidder's value for the two items can be summarized by a single parameter. Burguet (1999) shows that the A-RTC auction raises more revenue than the SA auction for general concave utility functions. Burguet (2002) identifies the seller's optimal auction when bidders are risk neutral. He also shows that the ascending right to choose auction in which the seller doesn't disclose which item is sold at the first stage raises more revenue than the conventional A-RTC auction. Finally, Burguet (2002) demonstrates that prices decline in successive stages of the A-RTC auction (see also Gale and Hausch 1994), as observed empirically (Ashenfelter and Genesove 1992).

Meneze and Monteiro (1998) characterize equilibrium in so-called "pooled" auctions. In these auctions n bidders simultaneously submit sealed bids. The highest bidder first chooses one of the k (possibly different) items, the second highest bidder then chooses one item from the remaining $k - 1$ items, and so on. Each of the k winners pay their own bid. Salmon and Iachini (2003) experimentally compare pooled auctions to simultaneous ascending auctions. They find that bids are above the level predicted by the risk-neutral bidding model. Indeed, bids are so high that subjects, on average, earn losses.

2. Design of the Experiment

We conducted two treatments as shown by the leftmost column of Table 1. We compared the simultaneous ascending (SA) auction with the ascending right-

TABLE 1. Experimental design

	Number of sessions	Number of subjects	Number of periods	Bidders per group	Value preferred item	Value inferior item
SA	1	8	30	4	$U[20, 920]$	20
A-RTC	2	16	15	4	$U[20, 920]$	20

to-choose (A-RTC) format. The experiments were conducted at Caltech where subjects were recruited, eight at a time, for an experiment lasting thirty periods. There were three groups of eight subjects in total and each group was divided into two groups of four bidders.³

The first group of eight subjects participated in fifteen periods of the A-RTC auction, followed by fifteen periods with a different auction format.⁴ The order was reversed for the second group of eight subjects. A final group of eight subjects, divided into two groups of four bidders, participated in thirty periods of the SA auction. Bidders' values were randomly determined once (as described below) for each of the thirty periods. This same set of values was used for each group of four bidders in both treatments, so that comparisons between auction institutions are direct. Written instructions were distributed and then read out loud to the subjects. The remainder of the experiment was computerized.

In each period of the experiment, two items, *A* and *B*, were auctioned among four bidders. The value each bidder assigned to each object was determined as follows. With probability one-half item *A* was designated as her "preferred item" and with probability one-half item *B* was designated as her "preferred item." A bidder's value for the preferred item was uniformly distributed between 20 and 920 points and her value for the "inferior" item was 20 points (see the final two columns in Table 1). Which item a bidder preferred, and her value for the preferred item, was independent across bidders and across periods. When a bidder won an item, she earned a number of points equal to her value for the item minus her bid.⁵ At the end of the experiment points were converted to dollars at a rate of 300 points = \$1. The experiments lasted between 1.5 and 2 hours and subjects made about \$20 on average.

In the SA auction, the items were simultaneously sold in two ascending bid auctions. A bidder could bid on at most one item at a time, that is, she could not bid on an item if she was the current high bidder on the other item. A bidder's

3. The composition of the group was fixed over the course of the experiment, but to avoid repeated game effects subjects were not informed of this fact.

4. The data from this part of the experiment are not relevant for this paper.

5. A bidder who didn't win either item received an "outside option" payoff of 80 points. The value distribution of $U[20, 920]$ followed from the outside option combined with values drawn from $U[100, 1000]$.

computer screen displayed a clock which counted down to zero and also displayed, for each item, the current high bid and the ID of the current high bidder. Whenever a new bid was submitted on either item that exceeded the current high bid on that item, the clock was reset. When the clock counted down all the way to zero, the current high bidder for each item won the item and paid the winning bid.

In the A-RTC auction, two ascending auctions were run sequentially, with a clock counting down to zero in both stages. The clock was reset whenever a new bid was submitted that exceeded the current high bid. When the clock counted down all the way to zero in the first auction, the current high bidder won the right to choose either item *A* or *B* and had to pay her bid. After the winner selected an item, the other bidders observed which item was chosen and the remaining item was sold by an ascending auction.

3. Theoretical Background

Our theoretical analysis closely follows Burguet (1999). Bidder *i*'s type is given by (θ_i, v_i) , where $\theta_i \in \{A, B\}$ indicates the item bidder *i* prefers and v_i denotes her private value for the preferred item. Bidders' values are identically and independently distributed: the value for the preferred item is assumed to be uniformly distributed on $[\underline{v}, \bar{v}]$ and the value for the inferior item is \underline{v} . Bidders can acquire at most one item. All bidders have a common utility function u . When bidder *i* wins an item of value v at price p then her utility is given by $u(v - p)$ where $u(0) = 0$, $u'(\cdot) > 0$, and $u''(\cdot) \leq 0$, that is, we assume bidders are (weakly) risk averse.

3.1. The SA Auction

The simultaneous ascending auction is the simplest format to analyze since risk aversion plays no role. It is an equilibrium strategy to bid on the item that yields the highest profit as long as this profit is positive. Hence the bidder with the overall highest value wins her preferred item, and the other item goes to the bidder with the highest value for that item. It is straightforward to determine the resulting expected revenue:⁶

$$R_{SA} = (57/40)\underline{v} + (23/40)\bar{v}.$$

6. There is a 1/8 chance that all four bidders prefer the same item in which case expected revenue is $((2/5)\underline{v} + (3/5)\bar{v}) + \underline{v}$, there is a 1/2 chance that three bidders prefer the same item in which case expected revenue is $((1/2)\underline{v} + (1/2)\bar{v}) + \underline{v}$, and there is a 3/8 chance that two bidders prefer one item and the other two bidders prefer the other item in which case expected revenue is $2((2/3)\underline{v} + (1/3)\bar{v})$.

3.2. The A-RTC Auction

Equilibrium in the ascending right-to-choose auction is derived using a backward induction argument. The second-stage auction is a standard English auction for the one remaining item, and it is a dominant strategy to bid up to one's value. Note that risk aversion plays no role in the second stage, and that there is always some bidding activity since the values of the three remaining bidders for the item left are at least $\underline{v} > 0$. Expected second-stage revenues are easily computed⁷ as $(17/20)\underline{v} + (3/20)\bar{v}$.

The symmetric equilibrium bidding functions of the first stage are defined recursively. We assume that bidders observe when others drop out and that exit is irrevocable. Consider the case where all four bidders are still active. Unlike the standard ascending auction, a bidder will not bid all the way up to her value. The reason is that her profit vanishes when the price level approaches her value and she strictly prefers dropping out and taking her chances in the second stage auction. To derive second-stage expected payoffs, note that the first bidder to drop out can only make a positive profit if all her rivals prefer the other item, which happens with probability $1/8$. So, for a bidder whose value for her preferred item is v , the utility of participating in the second-stage auction is 0 with probability $7/8$ and $u(v - \underline{v})$ with probability $1/8$. The certainty equivalent of this lottery is $u^{(-1)}((1/8)u(v - \underline{v}))$, where $u^{(-1)}(\cdot)$ denotes the inverse of the utility function. Hence it is optimal for this bidder to drop out of the first auction at price level

$$B(v) = v - u^{(-1)}((1/8)u(v - \underline{v})).$$

If the first bidder drops out at price level p_4 , say, the remaining bidders infer her value to be $v_4 = B^{-1}(p_4)$. The second-stage utility of the person who drops out next is therefore 0 with probability $3/4$, $u(v - \underline{v})$ with probability $1/8$, and $u(v - v_4)$ with probability $1/8$. Hence, it is optimal for the next bidder to drop out at the price level

$$B(v|v_4) = v - u^{(-1)}((1/8)u(v - \underline{v}) + (1/8)u(v - v_4)).$$

Finally, after a second bidder drops out at price level p_3 , the remaining two bidders infer this bidder's value to be $v_3 = B^{-1}(p_3|v_4)$. Of the remaining two bidders, the second-stage utility of the next bidder to drop out is 0 with probability $1/2$, $u(v - v_3)$ with probability $1/4$, $u(v - v_4)$ with probability $1/8$, and $u(v - \underline{v})$ with probability $1/8$. She drops when the price reaches

7. With probability $1/2$ the item is preferred by at most one bidder in which case revenue is \underline{v} . With probability $1/4$ the second-stage price is determined by the lowest-value bidder in which case revenue is $(4\underline{v} + \bar{v})/5$ and with probability $1/4$ the second-stage price is determined by the second-lowest-value bidder in which case revenue is $(3\underline{v} + 2\bar{v})/5$.

$$B(v|v_3, v_4) = v - u^{(-1)}(1/8)u(v - \underline{v}) + (1/8)u(v - v_4) + (1/4)u(v - v_3).$$

From the above expressions it is straightforward to determine the effects of risk aversion. Following standard terminology, we say that bidders with utility function $u_1(\cdot)$ are *more risk averse* than bidders with utility function $u_2(\cdot)$ if u_1 is a strictly concave transformation of u_2 , that is, $u_1(\cdot) = g(u_2(\cdot))$ with $g(0) = 0$ and $g(\cdot)$ strictly concave. Write $B_i(v)$, $B_i(v|v_4)$, and $B_i(v|v_3, v_4)$ for the bidding functions above where the utility function u is replaced with u_i . We say more risk averse bidders *bid more aggressively* if $B_1(v) > B_2(v)$, $B_1(v|v_4) > B_2(v|v_4)$, and $B_1(v|v_3, v_4) > B_2(v|v_3, v_4)$ for all $v > v_3 > v_4 > \underline{v}$.

PROPOSITION 1. *In the A-RTC auction, more risk averse bidders bid more aggressively.*

Proof. By including a term $u(0) = 0$, the three bidding functions above can be written as

$$v - u^{(-1)}\left(\sum_{k=0}^K p_k u(x_k)\right),$$

for $K = 1, 2, 3$, where $\sum_{k=0}^K p_k = 1$.⁸ Since $g^{(-1)}$ is strictly convex, Jensen’s inequality yields

$$u_1^{(-1)}\left(\sum_{k=0}^K p_k u_1(x_k)\right) = u_2^{(-1)}\left(g^{(-1)}\left(\sum_{k=0}^K p_k g(u_2(x_k))\right)\right) < u_2^{(-1)}\left(\sum_{k=0}^K p_k u_2(x_k)\right).$$

Hence, more risk averse bidders bid more aggressively. □

We next determine the effects of risk aversion on revenues. Let Y_k denote the k -th order statistic from four value draws. Total expected revenue of the A-RTC auction is given by

$$R_{ARTC} = E_{Y_2, Y_3, Y_4}(B(Y_2|Y_3, Y_4)) + (17/20)\underline{v} + (3/20)\bar{v}.$$

An easy corollary to Proposition 1 is that with risk averse bidders the A-RTC auction yields more revenue.

PROPOSITION 2. *The expected revenue of the A-RTC auction exceeds that of the SA auction when bidders are risk averse. The formats are revenue equivalent when bidders are risk neutral.*

8. Here $x_0 = 0$, $x_1 = v - \underline{v}$, $x_2 = v - v_4$, $x_3 = v - v_3$, and $p_1 = 1/8$, $p_2 = 1/8$, $p_3 = 1/4$, $p_0 = 1 - \sum_{k=1}^K p_k$ for $K = 1, 2, 3$.

TABLE 2. Predicted and observed per-round revenues

	SA	A-RTC
Observed Revenue: R_{obs}	579 (294)	691 (341)
Risk Neutral Prediction: R_{pred}^{RN}	556 (290)	553 (320)
Risk Averse Prediction: R_{pred}^{RA}	556 (290)	664 (345)
Efficiency	100.0%	98.4%
Bidders' Profits: $\pi_{bidders}$	479 (329)	356 (291)

Note: Standard deviations in parentheses.

Proof. In light of Proposition 1, the first claim follows from the second one since risk aversion has no effect in the SA auction while it raises bids in the A-RTC auction. When bidders are risk neutral, $u(x) = u^{(-1)}(x) = x$, so that $B(v|v_3, v_4) = (1/2)v + (1/8)\underline{v} + (1/8)v_4 + (1/4)v_3$. A straightforward computation shows that $R_{ARTC} = (57/40)\underline{v} + (23/40)\bar{v} = R_{SA}$ in this case. \square

In the next sections we report experimental evidence on the revenue generating properties of the SA and A-RTC auctions.

4. Results: Aggregate Data

The first row of Table 2 gives the mean per-period observed revenue for each auction format. (Standard deviations are in parentheses.) The second row gives mean predicted revenue, conditional on the induced values actually used in the experiment, for risk-neutral bidders. Finally, the third row of the table shows predicted revenue for the risk aversion parameter estimated in Section 5.⁹ Note that observed revenues are in line with the theoretical results with risk averse bidding—the A-RTC auction generates more revenue than the SA. The standard deviation of revenues, however, is higher for the A-RTC auction. Figure 1 provides another way of looking at the revenues properties of the auction formats. The empirical c.d.f. of revenues in the A-RTC auction first-order stochastically dominates the same c.d.f. for the SA auction. Hence, a seller whose utility is increasing in revenue prefers the A-RTC auction to the SA auction.

We next discuss tests of whether the revenue differences between auctions are statistically significant. For each observation $i = 1, \dots, 60$ in the A-RTC auction, define $\delta_i^{ARTC} = R_{obs,i}^{ARTC} - R_{pred,i}^{ARTC}$ to be the difference between the i -th

9. The estimation procedure is discussed in the next section. Note that for the SA auction, predicted revenue is the same whether bidders are risk neutral or risk averse.

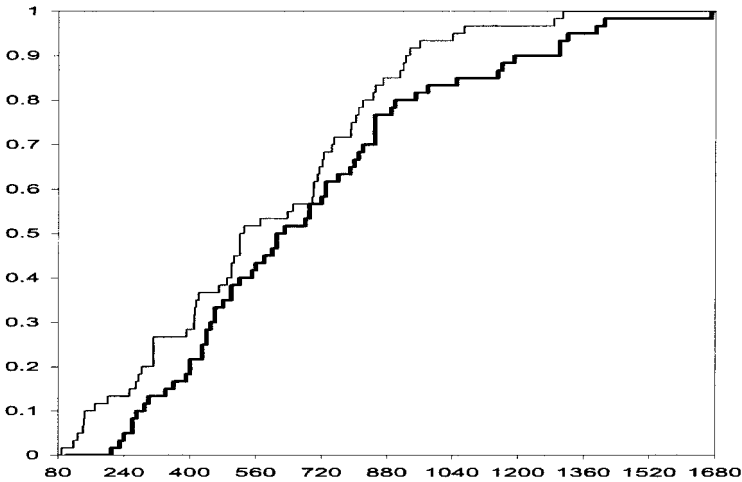


FIGURE 1. The cumulative distribution of observed revenues in the simultaneous ascending auction (thin black line) and ascending right-to-choose auction (heavy black line).

observed revenue and the i -th predicted revenue under the risk-neutral bidding model for the given induced values. Define δ^{SA} analogously for the SA auction. The means (standard deviation) of δ^{ARTC} and δ^{SA} are $\mu_{ARTC} = 138$ (107) and $\mu_{SA} = 23$ (64), respectively. Consider the null hypothesis that $\mu_{ARTC} = \mu_{SA}$ against the alternative that $\mu_{ARTC} > \mu_{SA}$. Using a t -test we can reject the null hypothesis at even the 1% significance level ($t = 7.15$ with 58 degrees of freedom).¹⁰ Hence, the additional revenue generated by the A-RTC auction (beyond the risk-neutral predicted amount) is significantly greater than the extra revenue generated by the SA auction.

The fourth row of Table 2 shows that both auction formats are highly efficient. Here efficiency is defined as follows: Denote realized surplus by $V_{realized}$, denote expected surplus were the items randomly allocated by V_{random} , and denote the efficient level of surplus by $V_{optimal}$. Efficiency is defined as

$$\frac{V_{realized} - V_{random}}{V_{optimal} - V_{random}}.$$

This measure of efficiency uses the surplus obtained by randomly allocating the items as a benchmark for measuring performance, and hence is more conservative than the standard efficiency measure of $V_{realized}/V_{optimal}$. The last row of Table 2 reports observed bidders' profits, and shows that the SA auction yielded higher expected profits, but also a higher variance.

10. See p. 435 of Mood, Graybill, and Boes (1974) for a description of this test. We can also reject the hypothesis that mean revenue in the SA auction is the same as mean revenue in the A-RTC auction. (The p -value is 0.058).

TABLE 3. Maximum likelihood estimates

Model	r	σ	$\log L$
Risk Neutral	0	107.8 (7.0)	-731.9
Risk Averse	0.61 (0.04)	69.9 (4.5)	-679.8

5. Results: Individual-Level Data

Here we use maximum likelihood techniques to estimate a common risk-aversion coefficient from the observed winning bids in the first stage of the A-RTC auction. The econometric model underlying the estimations is given by

$$b_i^{obs} = b_i^{pred}(r) + \varepsilon_i,$$

where ε_i is a normally distributed random variable with mean zero and variance σ^2 , b_i^{obs} is the observed bid, and $b_i^{pred}(r)$ is the predicted bid assuming a model of constant relative risk aversion (CRRA) with bidders' utilities given by $u(x) = x^{1-r}$. As highlighted by Proposition 1, predicted bids depend on the degree of risk aversion as measured by the CRRA coefficient, r . The log-likelihood corresponding to this model is given by

$$\log L = - \sum_{i=1}^n \left(\frac{b_i^{pred}(r) - b_i^{obs}}{\sigma} \right)^2 - n \log(\sqrt{2\pi}\sigma),$$

with n the number of observations used in the estimation procedure.

By maximizing the log-likelihood with respect to r and σ we obtain estimates for the risk-aversion model (see the second row of Table 3). The estimated risk aversion coefficient shows that bidders have the utility function $u(x) = x^{0.39}$. We compare the risk averse model with the risk-neutral model for which the risk aversion coefficient is fixed $r = 0$ (see Table 3). Obviously, the introduction of the single risk aversion coefficient results in a significant increase in likelihood.

It is interesting to use the parameter estimates of the risk aversion model to predict revenues for the values actually employed in the experiment. Figure 2 shows observed and predicted cumulative revenue distributions, both for the risk-averse and the risk-neutral model, for the two auction formats. The null hypothesis that the observed revenues are generated by the risk averse model cannot be rejected at the 10% significance level, using a standard Kolmogorov–Smirnov test.¹¹

11. The test statistic is 0.93 while the critical value for the 10% significance level is 1.22.

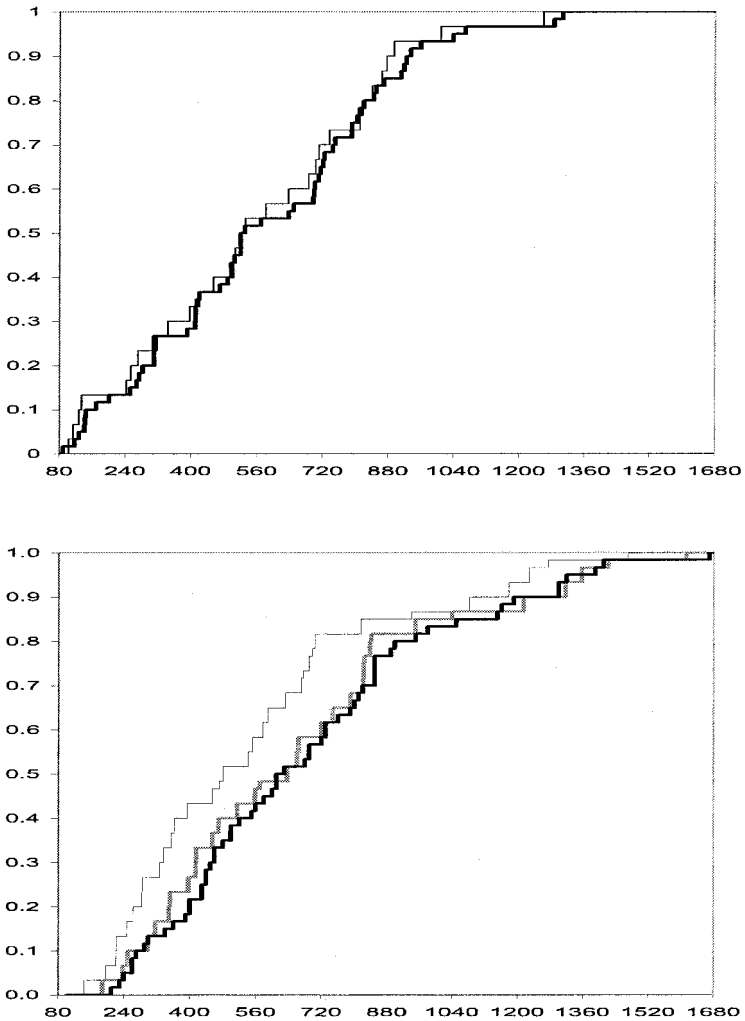


FIGURE 2. Predicted revenues. Risk neutral bidders (thin black lines), risk averse bidders (grey lines), and observed revenues (heavy black lines) in the SA auction (top) and A-RTC auction (bottom).

6. Conclusions

This paper develops a simple model of the ascending right to choose (A-RTC) auction. It shows that when bidders are risk averse, selling items via an ascending right to choose auction raises more revenue than would be obtained by selling the items in a simultaneous ascending (SA) auction. If bidders are risk neutral, both formats raise the same revenue. Our experimental results clearly

demonstrate that the ascending right to choose auction raises significantly more revenue than the simultaneous ascending auction.

One extension is to consider bidders who value more than one item. It is an open question whether the revenue superiority of the ascending right to choose auction extends to richer valuation structures where the simultaneous ascending auction has proven to perform well (e.g., Plott and Salmon 2004). It would also be interesting to model and experimentally assess the revenue properties of alternative right to choose auctions formats, and comparing their performance to the optimal mechanism.

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